## STICHTING MATHEMATISCH CENTRUM 2e BOERHAAVESTRAAT 49 AMSTERDAM

SM 87 ar

Wilcoxon's two sample test.

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## Receptuur

## Wilcoxon's two sample test

De toets van Wilcoxon voor twee steekproeven

Wilcoxon's two sample test may be used if one wants to investigate whether two samples are drawn from the same population. We first give an example of the problem of two samples.

Two different diets A and B are to be compared concerning their influence on the weight of rats; 11 rats are given diet A and 9 rats get diet B. At a certain moment the weight of each rat is determined. The results are (in gramms)

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A: 155, 154, 149, 163, 146, 150, 154, 161, 148, 145, 149, B: 156, 163, 153, 163, 165, 156, 151, 157, 167.
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On account of these observations one wants to investigate whether, in general, a difference exists between the weight of rats with diet A and the weight of rats with diet B, i.e. one wants to investigate whether the observed weights of the rats can be regarded as two samples drawn from the same population.

In general the problem may be formulated as follows.

Let  $x_1, \ldots, x_m$  be m independent observations of a random variable  $\underline{x}^1$ ), and let  $y_1, \ldots, y_n$  be n independent observations of a random variable  $\underline{y}$ , where  $\underline{x}$  and  $\underline{y}$  are distributed independently. Then the hypothesis  $H_0$  to be tested states that the probability distributions of  $\underline{x}$  and  $\underline{y}$  are identical.

The test to be described is a distribution free test, i.e. for the applicability of this test no assumptions are necessary on the form of the distributions of  $\underline{x}$  and  $\underline{y}$ . Thus e.g.  $\underline{x}$  and  $\underline{y}$  need not be normally distributed.

The test statistic of the Wilcoxon two sample test will be denoted by W and is calculated from the observations as follows. Each observation of  $\underline{x}$  is compared with each observation of  $\underline{y}$  and W is equal to twice the number of pairs of observations, for which the observation of  $\underline{x}$  is larger than the observation of  $\underline{y}$  plus (once) the number of pairs of observations for which the observation of  $\underline{x}$  equals the observation of  $y^2$ ).

It will be clear that  $\underline{W}$  assumes a small value if the observations of  $\underline{x}$  are

<sup>1)</sup> Random variables are distinguished from numbers (e.g. from the values they take in an experiment) by underlining their symbols.

<sup>&</sup>lt;sup>2</sup>) Usually the statistic  $U = \frac{1}{2}W$  is used for this test (cf [1]). In order to avoid fractions the statistic W is introduced in [2].

predominantly smaller than the observations of  $\underline{y}$  and a large value in the reverse situation.

The calculation of W for the example mentioned above is shown in scheme 1. 1) In this scheme the pooled samples of  $\underline{x}$  and  $\underline{y}$  are ranked according to increasing size (cf. column 1 and 2). Equal observations are placed on the same line.

Column 3 contains for each observation of  $\underline{x}$  the contribution to W, i.e. twice the number of observations of  $\underline{y}$  which are smaller than this observation of  $\underline{x}$  plus the number of observations of  $\underline{y}$  which are equal to it. Addition of the numbers in column 3 gives W = 34.

Column 4 contains the sizes of the ties, thus for each line the number of observations in the pooled samples on that line. Addition of the numbers in column 4 gives N = m + n = 20.

SCHEME I

Calculation of W

ľ	2	3	4	5
observa	tions of	contribution	sizes of	12
X	y	to W	ties (t)	T*
145		0	I	I
146 148		0	I	I
148		0	I	I
149,149		0	2	8
150		0	I	I
	151		r	I
	153		I	I
154,154		4,4	2	. 8
155		4	I	Ĭ
	156,156		2	8
	157		I	I
161		10	I	I
163	163,163	12	3	27
	163,163		I	I
	167		I	I

Column 5 contains the cubes of the sizes of the ties; their sum D is 62.

The distribution of  $\underline{W}$  assuming  $H_0$  to be true is known. This distribution is symmetric if no ties are present.  $\underline{W}$  then assumes even values and varies between 0 and 2 mn. The mean  $\mu$  and variance  $\sigma^2$  are given by (cf. [2] and [4])

<sup>1)</sup> An (analogous) scheme for cases with large samples and (or) large ties may be found in [2].

$$\mu=mn$$

(2) 
$$\sigma^2 = \frac{mn(N^3 - D)}{3N(N - I)} = \frac{1}{3}mn(N + I) - \frac{mn}{3N(N - I)}(D - N).$$

If one wants to test  $H_0$  against the alternative hypothesis that  $\underline{x}$  is systematically larger or smaller than  $\underline{y}$  a two-sided test is applied. The critical region of this two-sided test consists of large and small values of W.

It may occur that one wants to test  $H_0$  against the alternative hypothesis that  $\underline{x}$  is systematically smaller than  $\underline{y}$ . In this case a lower onesided critical region is used consisting of small values of W. If one wants to test  $H_0$  against the alternative hypothesis that  $\underline{x}$  is systematically larger than  $\underline{y}$  the upper onesided critical region is used consisting of large values of W.

Table 1 and 2 contain the lower critical values of W for the twosided test with  $\alpha = 0.1$  respectively  $\alpha = 0.05$  for  $m + n \le 40$  and  $m \le n$ . The upper critical values are found by substracting the tabulated value from 2 mn. The sample sizes m and n are interchangeable. The tables may also be used for the onesided test with  $\alpha = 0.05$  respectively  $\alpha = 0.025$ . These tables are taken from [2], where also tables are given of the exact tailprobabilities for  $m \le n \le 10$  and a table of the critical values with  $\alpha = 0.02$  (twosided) for  $m + n \le 40$ ,  $m \le n$  and  $n \ge 11$ . Strictly speaking these tables only hold for cases without ties, but they give a reasonable approximation for cases with small ties.

If m and n are large and if moreover the difference between m and n and the differences between the sizes of the ties are not too large,  $\underline{W}$  is approximately normally distributed with mean and variance according to (1) and (2). This fact may be used if m + n > 40 and also if  $m + n \leq 40$  and the sizes of the ties are too large to use the tables 1 and 2. In this case one calculates

$$u = \frac{|W - \mu| - I}{\sigma}$$

for the twosided test,

$$u = \frac{W - \mu + 1}{\sigma}$$

for the lower onesided test, and

$$u = \frac{W - \mu - I}{\sigma}$$

for the upper onesided test.

<sup>1)</sup> The term  $\pm$  1 in the numerator is the correction for continuity.

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	19	746 260 276 276
	18	518 246 272 272
d) 1)	17	167 207 247 208 247 209 209 209 209 209 209 209 209 209 209
(twosided) 1)	91	100 1100 1100 1100 1100 1100 1100 1100
= 0,1 (t	15	144 170 170 170 170 170 170 170 170 170
and α =	14	175 175 175 177 177 177 177 177 177 177
$n \leq m$	13	107 107 108 108 109 109 109 109 109 109 109 109 109 109
n ≤ 40,	12	8 6 1 1 7 7 1 1 7 8 8 1 1 7 7 8 8 1 1 7 8 8 1 1 7 8 8 1 1 7 8 8 1 1 7 8 8 1 1 7 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9
m + m	II	89 80 10 10 10 10 10 10 10 10 10 1
ABLE: test for	IO	50 47 88 80 11 10 10 10 10 10 10 10 10 10 10 10 10
T/ sample	6	44 49 45 45 60 1 1 1 2 1 2 1 2 1 2 1 2 1 2 1 2 1 2 1
's two	~	86 4 4 5 5 6 5 6 5 7 8 8 8 8 9 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1
o x o n's		22 6 4 8 7 5 6 6 6 7 7 8 8 6 6 6 1 1 1 1 2 1 2 1 1 1 1 1 2 1 1 1 1
Wilc	9	40048 686460369 8674888 8989 9811
for	7.	8 1 1 1 1 4 4 8 8 8 8 4 4 4 5 2 2 2 8 8 9 9 6 8 8 8 8 8 9 8 8 8 8 8 8 8 8 8
al values	4	4408014 080148 888 84444 888 889 898 888 888 888 888
Critical	3	00444088 00444088 04 44008 008 68 88 6444
	7	0 0 0 4 4 4 4 4 4 4 6 0 0 0 0 0 0 0 0 0
	<b></b>	

 $\mathfrak{m}$ 

0 0 

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20	. 254	
19	224 238	252
18	198 210 224	236 250 250
17	174 186 198 100	222 234 246 246
7 21	150 162 172 184 194	206 218 228 240
I 5	128 138 150 160 170	190 202 212 222 232 232
14	110 118 118 119 119 1196	$1/\infty$ $0/0$ $1/0$
13	100 100 100 100 100 100 100 100 100 100	160 170 178 186 196 204 214
I 2	74 82 90 117 130 138	146 154 162 178 194 202
}( }(	66 66 74 88 102 110 124	138 144 166 174 188 188
10	34 28 35 28 8 9 9 1 I	116 122 124 148 160 172 172
6	25 4 4 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5	100 106 1178 1188 1188 1190 1190
∞	3 % & 4 % 4 % 6 % 6 % 6 % 6 % 6 % 6 % 6 % 6 %	86 90 100 100 134 138 138
7	51 52 48 52 56 48 55 56 58 58 58 58 58 58 58 58 58 58 58 58 58	72 88 88 100 100 100 100 100 100 100 100 1
9	101 102 102 103 103 103 103 103 103 103 103 103 103	85 95 95 95 95 95 95 95 95 95 9
· .	4901491 817498 8488 9	4488833988888
*	$0.440 \times 0.1110 \times 0.2240 \times 0.000 \times 0.$	8 6 6 6 7 4 4 4 4 4 4 4 5 5 5 5 5 5 5 5 5 5 5 5
3	1 0 4 4 4 4 4 9 8 8 8 9 9 1 1 1 1 4 4 6	15 178 178 178 178 178 178 178 178 178 178
7	111000004444	2
· hand		footnote
#/ #/	24 50 50 11 12 12 12 12 12 12 12 12 12 12 12 12	21 22 23 24 25 26 29 30 31 33 34 35 36 37 37 37 37 37 37 37 37 37 37 37 37 37

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The tailprobabilities may then be found in a table of the normal distribution. If  $H_0$  is rejected one concludes that  $\underline{x}$  is systematically larger than  $\underline{y}$  if  $W > \mu$  and that  $\underline{x}$  is systematically smaller than  $\underline{y}$  if  $W < \mu$ .

In our example we have m = 11, n = 9, W = 34, D = 62. In table 2<sup>1</sup>) one finds for the lower critical value with  $\alpha = 0.05$  (twosided) 46; thus W being smaller than 46, the twosided tailprobability is smaller than 0.05.

The approximation with the normal distribution gives (cf. the tables 1 and 2 in [4])

$$\mu = 11 \times 9 = 99$$

and

$$\sigma^2 = 693 - \frac{8,684}{100}(62 - 20) = 689,35, \quad \sigma = 26,26.$$

Thus for the twosided case

$$u = \frac{|W - \mu| - I}{\sigma} = 2,44.$$

which gives a twosided tailprobability of 0,015.

If e.g. W = 250 in a case with m = 10, n = 15 and small ties, one first calculates (W being larger than  $\mu$ ) the upper critical value. For  $\alpha = 0.05$  (twosided) this critical value is (cf. table 2) 300-78 = 222. Thus W = 250 lies in the twosided critical region with  $\alpha = 0.05$ . More detailed data about this test may be found in [2].

## References

- [1] Mann, H. B. and D. R. Whitney, On a test of whether one of two random variables is stochastically larger than the other, Ann. Math. Stat. 18 (1947), 50-60.
- [2] Wabeke, Doraline en Constance van Eeden, Handleiding voor de toets van Wilcoxon, Report S 176 (M 65) of the Statistical Department of the Mathmatical Centre, Amsterdam, 1955.
- [3] Wilcoxon, F., Individual comparisons by ranking methods, Biometrics 1 (1945), 80—82.
- [4] Zaalberg, J., Auxiliary tables for Wilcoxon's two sample test, Statistica Neerlandica 12 (1958), 265—273.

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<sup>&</sup>lt;sup>1</sup>) The sample sizes m and n are interchangeable. Thus the critical values for m = 11, n = 9 are found at m = 9, n = 11.